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# Estimating Consumer Valuation of Earthquake Risk: Evidence from Japanese Housing Markets

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The relationships between seismic risk and rental and owner-occupied housing prices in the whole of Japan are examined. The empirical results from hedonic regressions with earthquake risk indices suggest that: (1) earthquake occurrence probability has a significantly negative effect on monthly housing rent, (2) the effect of earthquake probability seems to depend on the characteristics of the individual housing unit (e.g. age of dwelling) for owner-occupied housing, (3) the estimated risk premium is much larger for older buildings, and (4) the share of quake-resistant dwellings in the neighborhood area is significantly and positively related to the housing price of the individual unit. These results suggest that anti-seismic policies that target specific groups of dwellings, such as rental houses and older buildings, help to mitigate welfare loss due to earthquakes.

### Keywords

Earthquake; Hedonic price model; Risk premium

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## 1. Introduction

It is well known that Japan is one of the world's most earthquake-prone countries since it lies at the junction of four tectonic plates. According to the Opinion Survey on Disaster Prevention (Jiji Press, 2002), earthquakes (73.2%) are thought to be the most important risk factor among major natural and human disasters, such as fires (66.1%), floods (43.2%), and volcano eruptions (15.7%).

Since earthquakes are an exogenous risk factor which is tied to specific location, its risk premium should be capitalized into local housing and land prices. Estimating earthquake risk premium is important not only because it is the direct measure for the welfare loss due to earthquakes, but also because it is necessary for the evaluation of the effectiveness of anti-disaster policies.

Earthquake risk should be divided into two components: (1) exogenous occurrence probability, and (2) local attributes which amplify the damage of earthquakes. Since earthquake probability is purely exogenous and not under the policymaker's control, any policy instruments for disaster prevention should aim at minimizing earthquake damage. Since these two components are interrelated, i.e., anti-seismic policies may be extensively implemented in the region with high occurrence probability, omitting either of these components leads to incorrect results. For example, if we use occurrence probability as the index of earthquake risk, while we omit local attributes from the analysis, the impact of occurrence probability will be underestimated because anti-seismic policies are intensively implemented in the region with high occurrence probability. Therefore, we need to consider both of these components to assess the effectiveness of anti-seismic policies by using observational data.

In this paper, we will combine the household longitudinal data that cover all of Japan with seismic hazard information to estimate individual valuation of earthquake risk. Compared with previous studies, our contribution is as follows. First, we explicitly introduce several measures of earthquake risk into our analysis and distinguish their effects. As noted above, exogenous earthquake occurrence probability and damage-amplifying local attributes are used as the separate measures of earthquake risk. Secondly, compared with previous studies that focus on fairly small areas, we use nationwide longitudinal data in our analysis, which allow us to examine the entire effect of earthquake risk on the housing market in Japan. Thirdly, while previous studies mainly focus on land and rental markets, our dataset allows us to study a much wider range of the housing market in Japan. It provides detailed price information for both rental and owner-occupied housing: monthly rent, assessed values for property taxes and owner-provided values of owner-occupied housing.

Our empirical findings are as follows: (1) the earthquake occurrence probability has a significantly negative effect on the monthly housing rent, (2) the effect in the owner-occupied housing market is not as clear as that in the rental market; however, the effect seems to depend on the characteristics of the individual housing unit (e.g. age of dwelling), (3) the estimated risk premium is much larger for older buildings, (4) the share of quake-resistant dwellings in the neighborhood area is significantly and positively related to the housing price of the individual unit.

The paper is organized as follows. Section 2 briefly reviews the previous studies of earthquake risk in the housing market. Section 3 introduces the data used (Keio Household Panel Survey, KHPS) and explains the estimation method and variables. Section 4 presents the empirical results and interpretation. Section 5 summarizes the paper and presents some conclusions.

## 2. Previous Studies

In spite of its importance in disaster prevention policies, there have been only limited studies on the effect of earthquake risk on housing and land prices. Among others, Willis and Asgary (1997) evaluate the cost and benefit of anti-seismic policies by the contingent valuation method (CVM). Beron, Murdoch, Thayer, and Vijverberg (1997), introducing earthquake hazard indices as an additional source of variation, conduct a hedonic analysis of the residential housing prices in the San Francisco Bay area, and compare the estimated hedonic functions before and after the 1989 Loma Prieta earthquake. The results indicate that the hazard indices have a significantly negative impact on the housing prices in both time periods; however, their impact is greater in the pre-earthquake period, implying that the earthquake risk premium is overestimated before the Loma Prieta earthquake occurred. Naoi, Seko, and Sumita (2009a), using the same dataset as this paper, investigate whether homeowners and/or renters alter their subjective assessments of earthquake risks after an earthquake. They find that there are some modifications of individual assessments of earthquake risk following a major tectonic event, and that homeowners may initially underestimate earthquake risk in the pre-quake period. Brookshire, Thayer, Tschirhart, and Schulze (1985) examine the effects of the disclosure of a hazard map in California on land prices. It is found that the earthquake hazard indices have a significantly negative impact after the disclosure, but not before it.

The studies that are most closely related to ours in motivation are that by Nakagawa, Saito, and Yamaga (2007, 2009). While the former focuses on the rental market, the latter examines the impact on land market. Nakagawa et al. (2007) examine the impact of earthquake risk on housing rent by using an earthquake risk index taken from an earthquake hazard map compiled by the Tokyo Metropolitan Government with special reference to the new Building

Standard Law enacted in 1981. They find that housing rent is substantially lower in the areas with exposure to higher earthquake risk. Also, it is found that the rent of houses built prior to 1981 is discounted more substantially in risky areas than that of houses built after 1981. An important point to be noted is their use of listing prices rather than the actual rent paid. Although the use of listing prices of rental housing has several advantages, it will suffer from asymmetric information in a housing market, i.e., the seller has better information on the earthquake-resistant quality of the unit than the buyer, which might lead to biased estimates of earthquake risk premiums.

Nakagawa et al. (2009) empirically investigate the effect of earthquake risk on land prices, using the same earthquake risk index that was used by Nakagawa et al. (2007). Their result suggests that a higher earthquake risk is certainly related with lower land prices in each area.

### 3. Data and Methodology

#### 3.1 Data

The Keio Household Panel Survey (KHPS), sponsored by the Ministry of Education, Culture, Sports, Science and Technology, is the first comprehensive panel survey of households in Japan, conducted annually by Keio University since 2004. In wave 1, self-administered questionnaires were given to 4,005 male and female respondents, aged 20-69 years. These respondents were selected by stratified two-stage random sampling. If the primary respondent was married at the time of the survey, the same questionnaire was given to his/her spouse. The standard procedure for the KHPS was to send a pre-survey letter to the respondent and then provide a post-interview payment of 3,000 yen (approximately \$30) per household.

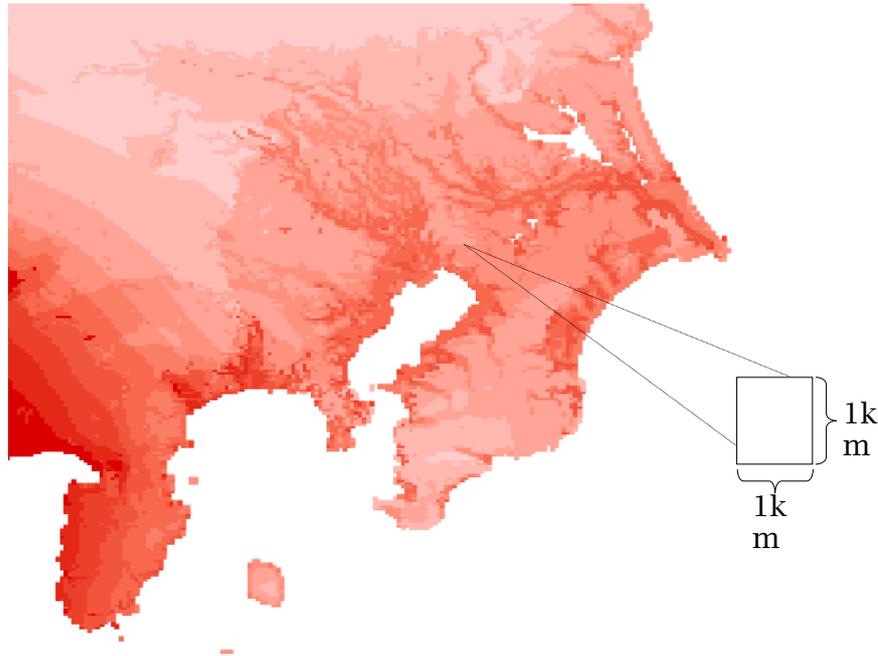
In the following analysis, three waves of the KHPS (2004–2006) are utilized to examine the relationship between seismic risk and housing prices in Japan, and to estimate the risk premium indices. As mentioned, various measures of housing prices are documented in the KHPS. For rental households, actual monthly rent paid is documented. For homeowners, assessed values for property taxes and owner-provided values of owner-occupied housing are documented.<sup>1</sup> The KHPS also provides detailed information on the type of housing; ownership status (owned, private rental, or public rental) and construction type (wooden or reinforced concrete building). Since risk premiums might critically depend on housing types, these information are necessary for evaluating the sole impact of seismic risk on the housing market, which are impossible in the previous studies due to data limitation.

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<sup>1</sup> The latter measure of housing price is constructed from the question about subjective assessment of the value of current residence (“How much do you think this lot/house would sell for on today's market?”).

The seismic risk measure is taken from the probabilistic seismic hazard map (PSHM) provided by the National Research Institute for Earth Science and Disaster Prevention (NIED).<sup>2</sup> The PSHM provides the probability of earthquake occurrence for a fixed time period and intensity. In the following analysis, we will use the occurrence probability of earthquakes with ground motions equal to or larger than the Japan Meteorological Agency (JMA) seismic intensity of 6 in the past 30 years, as our measure of seismic risk. The JMA seismic intensity scale, graded from 0 to 7, provides a measure of the strength of seismic motion.<sup>3</sup> An example of PSHM is shown in Figure 1.

**Figure 1** Example of the Probabilistic Seismic Hazard Map (PSHM)

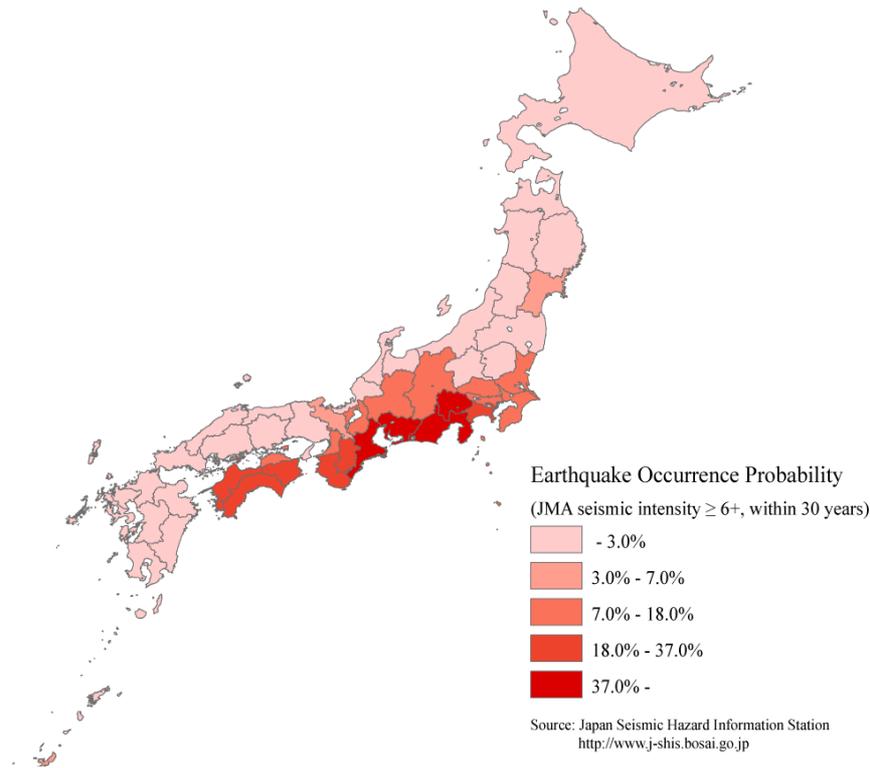


<sup>2</sup> The original data is available at <http://www.j-shis.bosai.go.jp/>.

<sup>3</sup> The JMA seismic intensity scale, which is measured with a seismic intensity meter, provides a measure of the strength of seismic motion. The typical situations and damages caused by an earthquake with a JMA seismic intensity of 6 are as follows: people have difficulty in trying to stand, wooden houses occasionally collapse, and walls and pillars may be damaged even for highly earthquake-resistant houses. For a full explanation of the JMA seismic intensity scale, see <http://www.jma.go.jp/jma/kishou/known/shindo/explane.html>.

Since the unit of observation in the original PSHM is defined based on the 3<sup>rd</sup> level mesh codes (1km meshed grid), city-level averages are calculated in order to match the seismic risk measures with the KHPS.<sup>4</sup> The resulting seismic risk measures are quite heterogeneous across prefectures; there are remarkably high earthquake probabilities in the southern coastal region (Figure 2). Moreover, these measures are highly diversified even within the same prefecture. Therefore, the seismic risk should be treated as a local attribute that is specific to fairly small areas (i.e. cities).

**Figure 2 Earthquake Occurrence Probability by Prefecture**



<sup>4</sup> This is because in the KHPS, the information on the respondent's location of residence is reported at the city/county-levels. The city-level averages of earthquake occurrence probabilities are calculated by ArcView 9.0.

While household perception toward seismic risk directly depends on the occurrence probabilities discussed above, it is also affected by the neighborhood characteristics of the residential region. Once an earthquake occurs, regions in which low quake-resistant dwellings are concentrated would suffer from immense damage. The city-level dwelling composition by its construction type is introduced to account for possible (negative) externalities. The data comes from the *2003 Housing and Land Survey of Japan* (Ministry of Land, Infrastructure and Transport, MLIT), which gives the fraction of dwellings with specific construction material; wooden, fire-proof wooden, reinforced steel-framed concrete, and steel-framed dwellings, for every city in Japan. Generally, wooden dwellings are thought to have the least quake-resistant quality. Regions crowded with these dwellings will have higher earthquake risk not only because wooden buildings can easily collapse, but also these buildings will be the major cause of fires after the earthquake.

### 3.1 Empirical Model and Variables

Our primary interest is on estimating the seismic risk premium. The hedonic regression model is given as follows:

$$p_i^{(\lambda)} = \alpha + \beta EQ_i + x_i \gamma + \varepsilon_i, \quad (1)$$

where  $p_i$  is the appropriate housing price measure for unit  $i$  (which will be actual rent, assessed values for property taxes or owner-provided house values, depending on the model to be estimated),  $EQ_i$  is the seismic risk measure (i.e. earthquake probability),  $x_i$  is the relevant set of explanatory variables, and  $\alpha$ ,  $\beta$ , and  $\gamma$  are parameters to be estimated. The Box-Cox transformation with parameter  $\lambda$  yields:

$$p_i^{(\lambda)} = \frac{p_i^\lambda - 1}{\lambda}, \quad (2)$$

The model becomes linear with  $\lambda = 1$  and semi-logarithmic with  $\lambda = 0$  as the special case. Following previous studies on hedonic analyses of the housing market,  $x_i$  includes the basic housing characteristics, such as number of rooms, floor and garden space, years since the unit was built, number of floors, and the time distance to the nearest station/bus stop. In addition to these basic characteristics, we also control dummies for construction type and ownership status of the dwelling, city size, and the region in which the unit is located. The definition and summary statistics of the variables are shown in Table 1.

**Table 1** Variable Definitions and Summary Statistics

Sample	Renter Households		Homeowners (detached)		Homeowners (detached)		Homeowners (condo)		Homeowners (condo)	
	Mean	(S.D.)	Mean	(S.D.)	Mean	(S.D.)	Mean	(S.D.)	Mean	(S.D.)
<i>Housing Prices</i>										
Housing rent (10,000 yen / month)	6.128	(3.386)	—	—	—	—	—	—	—	—
Market price (10,000 yen)	—	—	3229.0	(4831.1)	—	—	1777.2	(1675.0)	—	—
Assessed value (10,000 yen)	—	—	—	—	1372.2	(2258.3)	—	—	707.7	(1729.4)
<i>Earthquake Risk Indices</i>										
Earthquake occurrence probability	0.150	(0.162)	0.161	(0.195)	0.164	(0.201)	0.143	(0.114)	0.141	(0.119)
Neighborhood dwelling composition										
% wooden	24.261	—	30.182	—	30.226	—	19.449	—	19.509	—
% fire-proofed wooden	28.875	(13.237)	31.526	(14.768)	31.949	(15.009)	31.146	(12.198)	30.956	(12.813)
% steel-framed concrete	7.256	(3.026)	7.015	(3.286)	7.043	(3.324)	7.346	(2.697)	7.270	(2.853)
% concrete	39.351	(15.257)	30.955	(15.396)	30.444	(15.242)	41.897	(12.870)	42.102	(13.164)
% other types	0.257	(0.570)	0.322	(0.682)	0.337	(0.745)	0.161	(0.140)	0.163	(0.144)
<i>Dwelling Characteristics</i>										
Age of the building (years since built)	19.125	(13.012)	20.603	(14.789)	21.297	(15.007)	16.367	(9.180)	17.337	(8.881)
Number of rooms	3.358	(1.171)	6.285	(1.910)	6.388	(1.979)	4.278	(0.823)	4.285	(0.822)
Time-distance from the nearest	8.674	(7.163)	10.156	(9.617)	10.354	(9.816)	8.083	(6.488)	7.958	(6.313)
Number of stories of the building	3.072	(2.864)	1.915	(0.448)	1.908	(0.458)	7.040	(3.544)	6.924	(3.560)
Floor in which the room is located	1.942	(2.007)	—	—	—	—	3.294	(2.309)	3.243	(2.334)
Garden space	—	—	79.859	(112.690)	84.877	(117.098)	—	—	—	—
<i>N</i>	1,577		2,665		2,168		551		383	

Since it is well-known that the Wald statistics for the estimated coefficients of the right-hand-side variables are not invariant to changes in the scale of the transformed dependent variable (Spitzer, 1984; Davidson and MacKinnon, 1993), we instead perform and report the likelihood-ratio tests for each coefficient.

Given the estimated coefficients  $(\hat{\alpha}, \hat{\beta}, \hat{\gamma}$  and  $\hat{\lambda})$ , the fitted values and marginal effects are given as follows:

$$\text{Fitted Value: } \hat{p}(EQ, x) = \int [\hat{\lambda}(\hat{\alpha} + \hat{\beta}EQ + x\hat{\gamma} + \varepsilon) + 1]^{\frac{1}{\hat{\lambda}}} d\hat{F}(\varepsilon), \quad (3)$$

$$\text{Marginal Effect: } \hat{m}(EQ, x) = \hat{\beta} \int [\hat{\lambda}(\hat{\alpha} + \hat{\beta}EQ + x\hat{\gamma} + \varepsilon) + 1]^{\frac{1-\hat{\lambda}}{\hat{\lambda}}} d\hat{F}(\varepsilon), \quad (4)$$

where  $\hat{F}$  is an estimate of the true error distribution  $F$ . Following Abrevaya (2002), the “smearing” technique, which uses estimated residuals to approximate the error distribution, is used to obtain the estimates of  $\hat{p}$  and  $\hat{m}$ . In the following analysis, the marginal effect of earthquake probability is evaluated at a sample mean, i.e.  $\hat{m}(\overline{EQ}, \bar{x})$ . The earthquake risk premium is estimated by the changes in the fitted values of housing prices from an as-if situation (i.e. zero earthquake probability),  $\hat{p}(\overline{EQ}, \bar{x}) - \hat{p}(0, \bar{x})$ .

#### 4. Empirical Results

In the following analyses, we have split the sample into three groups based on the ownership status of the unit; rental houses, owner-occupied detached houses, and owner-occupied condominium units, and estimate equation (1) for each of these three groups.

##### 4.1 Baseline Result

Our baseline result is shown in Table 2. Five models are estimated for housing rent (Model [1]), owner-provided values and assessed values for property taxes of detached houses (Models [2] and [3]), and those of condominium units (Models [4] and [5]). In the table, estimated coefficients and marginal effects of seismic risk indices, i.e., earthquake occurrence probability and neighborhood dwelling composition, are reported.<sup>5</sup>

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<sup>5</sup> A set of dwelling characteristics and dummy variables for regions and survey years is also controlled, but omitted from the results. Dwelling characteristics included are as follows: age of the dwelling (years since built), number of rooms, number of stories of the building, time-distance from the nearest railway station/bus stop, garden space (for detached houses), floor in which the room is located (for condominium units), and dummies for the type of dwelling. The complete results are available upon request.

The results indicate that the earthquake occurrence probability has a significantly negative effect on housing rent and assessed values of detached houses, but not on other housing price measures.

The negative estimated coefficient and marginal effect in the housing rent model is consistent with previous studies (Naoi, Sumita, and Seko, 2007; Nakagawa et al., 2007). Our index of the earthquake risk premium indicates that the change in earthquake probability from a hypothetical riskless situation to the actual average level (i.e.  $0 \rightarrow 0.150$ ) leads to a 3,654 yen decrease in monthly rent, which implies that the earthquake risk premium accounts for approximately 6% of the average monthly rent ( $3,654 / 61.280 = 5.96\%$ ).<sup>6</sup>

As for detached houses (Models [2] and [3]), a significantly negative coefficient of earthquake probability is estimated for assessed values for property taxes, while it is not significant for owner-provided house values. This discrepancy will be further investigated in later sections. The earthquake risk premium index for Model [3] becomes roughly two million yen<sup>7</sup>, about 14.5% of the average house value.

The results for condominium units (Models [4] and [5]) suggest that the earthquake probability does not have any significant impact on their pricing. Unfortunately, this can be partly attributed to the limited sample sizes. As these units are concentrated in urban areas, perhaps regional earthquake occurrence probability does not have enough variation to estimate its true effect. Therefore, although the effects are estimated to be insignificant, further investigation might be required in future research.

As for the neighborhood dwelling composition, estimated coefficients become generally positive and are mostly significant. Since the wooden building is considered to have lower quake-resistance quality than other types of buildings, the result indicates that replacing wooden dwellings with other types of quake-resistant buildings leads to higher housing prices in each region.

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<sup>6</sup> Nakagawa et al.,(2007) report that the risk premium is about 3–6% of the housing value.

<sup>7</sup> The actual estimate is -199.75 (in 10,000 yen).

**Table 2 Effect of Seismic Risk Measures on Housing Prices**

Model Dependent Variable Sample Used	[1]			[2]			[3]			[4]			[5]		
	Rent (10,000 yen / month) Renter Households			Owner-Provided Values (10,000 yen) Homeowners (Detached House)			Assessed Values (10,000 yen) Homeowners (Detached House)			Owner-Provided Values (10,000 yen) Homeowners (Condominium)			Assessed Values (10,000 yen) Homeowners (Condominium)		
	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect
Earthquake Occurrence Probability	-0.7977	8.258 **	-3.2221	-0.7229	1.177	-12.7197	-7.2166	9.311 **	-95.3066	0.1343	0.124	7.6637	-1.6191	0.306	-24.2084
Neighborhood Dwelling Composition	(reference)			(reference)			(reference)			(reference)			(reference)		
% wooden	(reference)			(reference)			(reference)			(reference)			(reference)		
% fire-proofed wooden	0.0074	3.750 +	0.0301	-0.0015	0.030	-0.0268	0.0557	3.017 +	0.7357	0.0084	2.356 #	0.4784	0.0389	0.792	0.5815
% steel-framed concrete	0.0214	5.480 *	0.0865	0.1088	14.931 **	1.9138	0.3225	9.372 **	4.2589	0.0074	0.307	0.4234	-0.1104	1.165	-1.6505
% concrete	0.0167	35.748 **	0.0675	0.0479	64.241 **	0.8431	0.0526	5.830 *	0.6941	0.0141	9.257 **	0.8039	0.0377	1.071	0.5640
% other types	0.0868	5.476 *	0.3508	-0.2833	9.295 **	-4.9854	-0.1277	0.160	-1.6859	-0.2475	1.252	-14.1261	-0.8235	0.207	-12.3133
$\lambda$	0.3686	(0.0252) **		0.1736	(0.0124) **		0.2727	(0.0107) **		0.0202	(0.0344)		0.1136	(0.0283) **	
N	1,577			2,665			2,168			551			383		
Log likelihood	-3325.105			-23437.810			-17386.902			-4376.736			-2724.682		

Notes:

\*\*\*, \*, and + indicate that the estimated coefficient is significant at the 0.01, 0.05, and 0.10 levels, respectively. A set of dwelling characteristics and dummy variables for regions and survey years is also controlled but is omitted from the results. For Box-Cox transformation parameter ( $\lambda$ ), standard errors are reported instead of likelihood ratio test statistics ( $\chi^2(1)$ ).

#### 4.2 Changes in Effect of Earthquake Risk over Time

Interaction terms of earthquake occurrence probability with survey year dummies are introduced which take into account the fact that their effect may vary over time. The results are summarized in Table 3.

The overall results are similar to those reported in Table 1. However, in Model [1], the effect that earthquake probability has on housing rent substantially varies over time. The negative effect is the largest in 2006 and smallest in 2005. Given that the quality of typical housing is unchanged during our sample period, a possible interpretation for this result could be that household perceptions were updated for seismic risk. In November 2005, the Ministry of Land, Infrastructure and Transport announced a scandal where several structural designers had fabricated quake-resistance data in designs for condominiums and hotels in Tokyo, Chiba and Kanagawa prefectures, and that some of them might collapse in an earthquake with a JMA intensity of 5<sup>-</sup>.<sup>8</sup> Since the majority of rented units are condominiums and the fabrication took place for this type of building, we think that these updates seem to be prominent in the housing rent model.<sup>9</sup> Another possibility is that changes in the earthquake insurance market would affect household perception toward seismic risk. We find, however, that major revisions in the insurance premium policies took place in October 2007, which is out of our sample period. See Naoi, Seko, and Sumita (2009b) for details of the Japanese earthquake insurance market.

#### 4.3 Changes in the Effect of Earthquake Risk by Age of Building

We also include an interaction term between the earthquake occurrence probability and the dummy variables of the age of the building, given the possibility that the impact of earthquake risk on housing prices may depend on earthquake-resistant quality. Table 4 presents the results.

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<sup>8</sup> Under the 1981 Building Standard Law regulation, buildings must be strong enough to resist a quake with a JMA intensity of 6<sup>+</sup>.

<sup>9</sup> There are, however, several other events that can affect household perception toward seismic risk. For example, massive earthquakes, such as the Mid Niigata Pref. Earthquake in October 2004 (JMA intensity 7), Eastern Fukuoka Pref. Earthquake in March 2005 (JMA intensity 6<sup>-</sup>), and Miyagi Pref. Earthquake in August 2005 (JMA intensity of 6<sup>-</sup>), had taken place during our sample period.

**Table 3: Effect of Seismic Risk Measures on Housing Prices — Interacted with Survey Year Dummies**

Model Dependent Variable Sample Used	[1]			[2]			[3]			[4]			[5]		
	Rent (10,000 yen / month) Renter Households			Owner-Provided Values (10,000 yen) Homeowners (Detached House)			Assessed Values (10,000 yen) Homeowners (Detached House)			Owner-Provided Values (10,000 yen) Homeowners (Condominium)			Assessed Values (10,000 yen) Homeowners (Condominium)		
	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect
Earthquake Occurrence Probability $\times$ Survey Year Dummy															
2004	-0.8131	7.353 **	-3.2844	-0.9979	1.715	-17.5580	-9.0334	11.418 **	-119.277	-0.3812	0.694	-23.1953	-1.2089	0.118	-18.1438
2005	-0.6987	4.291 *	-2.8224	-0.3219	0.173	-5.6643	-8.2690	8.771 **	-109.184	0.2154	0.206	13.1075	-3.5846	1.027	-53.8017
2006	-0.8407	6.619 *	-3.3960	-0.8248	1.156	-14.5129	-4.4049	2.600	-58.1622	0.7825	2.388	47.6113	1.0171	0.066	15.2657
Neighborhood Dwelling Composition															
% wooden	(reference)			(reference)			(reference)			(reference)			(reference)		
% fire-proofed wooden	0.0075	3.766 +	0.0302	-0.0015	0.027	-0.0255	0.0552	2.949 +	0.7284	0.0079	2.129	0.4777	0.0399	0.842	0.5991
% steel-framed concrete	0.0214	5.478 *	0.0865	0.1095	15.131 **	1.9272	0.3260	9.523 **	4.3040	0.0084	0.407	0.5125	-0.1084	1.137	-1.6273
% concrete	0.0167	35.648 **	0.0675	0.0480	64.337 **	0.8447	0.0504	5.342 *	0.6660	0.0136	8.836 **	0.8250	0.0375	1.070	0.5634
% other types	0.0868	5.467 *	0.3507	-0.2810	9.136 **	-4.9442	-0.1208	0.143	-1.5951	-0.2708	1.541	-16.4780	-0.8694	0.233	-13.0485
$\lambda$	0.3686	(0.0252) **		0.1736	(0.0124) **		0.2731	(0.0107) **		0.0188	(0.0344)		0.1128	(0.0283) **	
<i>N</i>	1577			2665			2168			551			383		
Log likelihood	-3324.963			-23437.233			-17384.706			-4374.121			-2724.002		

Notes:

\*\* , \* , and + indicate that the estimated coefficient is significant at the 0.01, 0.05, and 0.10 levels, respectively. A set of dwelling characteristics and dummy variables for regions and survey years is also controlled but is omitted from the results. For Box-Cox transformation parameter ( $\lambda$ ), standard errors are reported instead of likelihood ratio test statistics ( $\chi^2(1)$ ).

**Table 4** Effect of Seismic Risk Measures on Housing Prices — Interacted with Age of the Building

Model Dependent Variable Sample Used	[1]			[2]			[3]			[4]			[5]		
	Rent (10,000 yen / month) Renter Households			Owner-Provided Values (10,000 yen) Homeowners (Detached House)			Assessed Values (10,000 yen) Homeowners (Detached House)			Owner-Provided Values (10,000 yen) Homeowners (Condominium)			Assessed Values (10,000 yen) Homeowners (Condominium)		
	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect	Coef.	$\chi^2(1)$	Marginal Effect
Earthquake Occurrence Probability $\times$ Age of the Building (Years since Built)															
Years $\leq$ 5	0.1795	0.149	0.7268	-1.5833	2.650	-27.8046	-9.3332	6.151 *	-123.147	0.2750	0.253	17.4802	-5.5763	1.514	-81.3381
5 < Years $\leq$ 10	-0.2558	0.511	-1.0355	-1.3750	2.456	-24.1473	-8.4724	6.871 **	-111.790	-0.3856	0.461	-24.5122	-1.7984	0.131	-26.2325
10 < Years $\leq$ 15	-0.4229	1.142	-1.7124	-1.1440	1.324	-20.0911	-9.6870	7.875 **	-127.816	-0.0454	0.004	-2.8853	-1.7164	0.093	-25.0354
15 < Years $\leq$ 20	-0.7299	3.635 +	-2.9550	-2.6033	7.079 **	-45.7181	-9.2926	7.325 **	-122.612	-0.3592	0.287	-22.8318	0.2197	0.002	3.2047
Years > 20	-1.3286	18.753 **	-5.3792	0.3813	0.244	6.6956	-5.5660	4.436 *	-73.4409	1.7672	9.543 **	112.339	0.8556	0.036	12.4804
Neighborhood Dwelling Composition															
% wooden	(reference)			(reference)			(reference)			(reference)			(reference)		
% fire-proofed wooden	0.0090	5.481 *	0.0365	-0.0002	0.000	-0.0033	0.0594	3.368 +	0.7836	0.0111	4.215 *	0.7077	0.0173	0.140	0.2528
% steel-framed concrete	0.0231	6.389 *	0.0933	0.1252	19.088 **	2.1986	0.3464	10.612 **	4.5705	0.0172	1.684	1.0945	-0.1169	1.194	-1.7056
% concrete	0.0162	33.446 **	0.0654	0.0461	57.327 **	0.8097	0.0478	4.726 *	0.6309	0.0138	9.211 **	0.8762	0.0291	0.579	0.4238
% other types	0.0314	0.733	0.1273	-0.2531	7.210 **	-4.4451	-0.0462	0.021	-0.6101	-0.3346	2.341	-21.2686	-1.3665	0.514	-19.9322
$\lambda$	0.3676	(0.0251) **		0.1768	(0.0122) **		0.2734	(0.0107) **		0.0179	(0.0341)		0.1193	(0.0283) **	
<i>N</i>	1,577			2,665			2,168			551			383		
Log likelihood	-3319.820			-23405.292			-17388.130			-4362.689			-2718.834		

Notes:

\*\*\*, \*, and + indicate that the estimated coefficient is significant at the 0.01, 0.05, and 0.10 levels, respectively. A set of dwelling characteristics and dummy variables for regions and survey years is also controlled but is omitted from the results. For Box-Cox transformation parameter ( $\lambda$ ), standard errors are reported instead of likelihood ratio test statistics ( $\chi^2(1)$ ).

It is found that the effect of earthquake probability substantially depends on the age of the individual housing unit and that in general, the estimated risk premium is much larger for older buildings. As for rental housing, the negative effect of the earthquake probability is the largest for a unit that is 20 years or older, and the effect becomes insignificant for relatively new units (ages 15 or less). Similar results, albeit to a lesser extent, can be observed for owner-occupied detached housing. The owner-provided value of detached housing, for which we cannot observe the effect of earthquake probability as a whole (see Table 2), is negatively influenced by the earthquake probability when the age of the unit is 15 – 20 years old.

## 5. Conclusion

The purpose of this paper is to examine the relationships between seismic risk, and rental and owner-occupied housing prices in Japan. The earthquake risk premium is estimated using hedonic price models based on the household longitudinal data that cover all of Japan.

Since earthquake risk is compounded of both the probability of an occurrence and the resulting damage to be expected, we have introduced two separate components of earthquake risk; exogenous earthquake occurrence probability and neighborhood dwelling composition, as the separate measures of earthquake risk, into our analysis. The results from hedonic regressions provide the following empirical findings: (1) the earthquake occurrence probability has a significantly negative effect on the monthly housing rent, (2) the effect in the owner-occupied housing market is not as clear as the rental market; however, the effect seems to depend on the characteristics of the individual housing unit (e.g. age of dwelling), (3) the estimated risk premium is much larger for older buildings, and (4) the share of quake-resistant dwellings in the neighborhood area is significantly and positively related to the housing price of the individual unit.

The result where the earthquake occurrence probability is shown to have a negative impact on housing rent, but not on owner-occupied housing values, partially mirrors the fact that quake-resistant quality is much lower in rental houses. This suggests that seismic retrofitting for rental housing might be an effective policy device for compensating earthquake risks. Also, given that the estimated risk premium is much larger for older buildings, policies that aim at the enhancement of seismic safety for such buildings might be an effective way to mitigate the welfare loss caused by earthquake risk. Furthermore, our result suggests that city-level dwelling composition has a large (negative) externality to the neighborhood dwellings, implying that, for example, an urban redevelopment project for congested wooden dwelling areas will be beneficial not only to the individual unit, but also neighborhood dwellings.

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